A Generalization of the I-MMSE Formula

Jian Song (joint work with Guangyue Han)

Department of Mathematics Department of Statistics & Actuarial Science Hong Kong University txjsong@hku.hk

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Problem and background

Comparison between old approach and our observation New proofs based on our observation Main results

The mutual information between two random elements X and Y, a measure of their mutual dependence, can be defined as

$$\begin{aligned} (X;Y) &= H(X) - H(X|Y) \\ &= \mathbb{E}\left[\log\frac{d\mu_{XY}}{d\mu_X \times \mu_Y}(X,Y)\right] \\ &= \int \int \log\frac{f_{XY}(x,y)}{f_X(x)f_Y(y)}f_{XY}(x,y)dxdy \end{aligned}$$

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Problem and background een old approach and our observation

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• Consider the discrete-time Gaussian channel

$$\mathbf{Y}^{\rho} = \rho \mathbf{X} + \mathbf{Z},$$

where Y^{ρ} : output, X: input, Z: Gaussian r.v., and ρ^2 : signal-to-noise ratio.

 In the paper Mutual information and minimum mean-square error in Gaussian channels by Guo, Shamai, and Verdu, the I-MMSE formula was obtained:

$$rac{d}{d
ho} I(X; Y^
ho) =
ho \mathbb{E} \left[(X - \mathbb{E}[X|Y^
ho])^2
ight].$$

Problem and background

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For the continuous-time channel

$$\mathbf{Y}_t^{
ho} =
ho \int_0^t \mathbf{X}_{m{s}} dm{s} + m{W}_t, \ t \in [0, T],$$

where W is a Brownian motion, the I-MMSE formula is

$$rac{d}{d
ho} I(X; Y^{
ho}) =
ho \mathbb{E} \int_0^T (X_s - \mathbb{E}[X_s|Y^{
ho}])^2 ds.$$

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We refer to the Shannon Lecture From Constrained Signaling to Network Interference Via An Information-Estimation Perspective given by *Shamai* for an overview of the applications of I-MMSE.

Consider the channel

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 $Y^{\rho} = \rho X + Z.$

$$I(X; Y^{\rho}) = H(Y^{\rho}) - H(Y^{\rho}|X) = H(Y^{\rho}) - H(Z).$$

$$\frac{d}{d\rho}I(X;Y^{\rho})=\frac{d}{d\rho}H(Y^{\rho})=-\frac{d}{d\rho}\mathbb{E}[\log f_{Y^{\rho}}(Y^{\rho})].$$

Recall that $Y^{\rho} = \rho X + Z$.

Computation of

$$rac{d}{d
ho}\mathbb{E}[\log f_{Y^{
ho}}(Y^{
ho})].$$

• The approach by Guo, Shamai and Verdu,

$$\frac{d}{d\rho}\int f_{Y^{\rho}}(y)\log f_{Y^{\rho}}(y)dy=\int \frac{\partial}{\partial\rho}f_{Y^{\rho}}(y)\log f_{Y^{\rho}}(y)dy.$$

• Our approach

$$\frac{d}{d\rho}\int\int f_X(x)f_Z(z)\log f_{XZ}(\rho x+z)dxdz = \mathbb{E}\left(\frac{d}{d\rho}\log f_{Y\rho}(Y^{\rho})\right).$$

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I-MMSE De Bruijn's identity

• Suppose Z now is a general random variable with density function $e^{T(z)}$, and $\mathbb{E}(X) = 0$. Consider the channel

$$\mathbf{Y}^{\rho} = \rho \mathbf{X} + \mathbf{Z}.$$

- It has been studied in Additive non-Gaussian noise channels: mutual information and conditional mean estimation by Guo, Shamai and Verdu.
- Note that $f_{Y^{\rho}|X=x}(y;x) = e^{T(y-\rho x)}, f_Y(y) = \int_{\mathbb{R}} e^{T(y-\rho x)} d\mu_X(x).$

$$\frac{d}{d\rho}f_{Y^{\rho}}(Y^{\rho}) = \frac{d}{d\rho}\int_{\mathbb{R}} e^{T(Y^{\rho} - \rho X)} \mu_X(dX)$$
$$= \frac{d}{d\rho}\int_{\mathbb{R}} e^{T(\rho(X - X) + Z)} \mu_X(dX)$$

$$=\int_{\mathbb{R}} e^{T(\rho(X-x)+Z)} T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= \int_{\mathbb{R}} e^{T(\rho(X-x)+Z)} T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$
$$= \int_{\mathbb{R}} f_{Y^{\rho}|X=x}(Y^{\rho}; x) T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= \int_{\mathbb{R}} e^{T(\rho(X-x)+Z)} T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= \int_{\mathbb{R}} f_{Y^{\rho}|X=x}(Y^{\rho};x) T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= f_{Y^{\rho}}(Y^{\rho}) \int_{\mathbb{R}} T'(Y^{\rho}-\rho x)(X-x))\mu_{X|Y^{\rho}}(dx;Y^{\rho}) \text{ (explain later)}$$

$$= \int_{\mathbb{R}} e^{T(\rho(X-x)+Z)} T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= \int_{\mathbb{R}} f_{Y^{\rho}|X=x}(Y^{\rho}; x) T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= f_{Y^{\rho}}(Y^{\rho}) \int_{\mathbb{R}} T'(Y^{\rho}-\rho x)(X-x))\mu_{X|Y^{\rho}}(dx; Y^{\rho}) \text{ (explain later)}$$

$$= f_{Y^{\rho}}(Y^{\rho}) \left(X \mathbb{E}[T'(Y^{\rho}-\rho X)|Y^{\rho}] - \mathbb{E}[T'(Y^{\rho}-\rho X)X|Y^{\rho}]\right)$$

I-MMSE De Bruijn's identity

$$= \int_{\mathbb{R}} e^{T(\rho(X-x)+Z)} T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= \int_{\mathbb{R}} f_{Y^{\rho}|X=x}(Y^{\rho};x) T'(\rho(X-x)+Z)(X-x)\mu_X(dx)$$

$$= f_{Y^{\rho}}(Y^{\rho}) \int_{\mathbb{R}} T'(Y^{\rho}-\rho x)(X-x))\mu_{X|Y^{\rho}}(dx;Y^{\rho}) \text{ (explain later)}$$

$$= f_{Y^{\rho}}(Y^{\rho}) \left(X \mathbb{E}[T'(Y^{\rho}-\rho X)|Y^{\rho}] - \mathbb{E}[T'(Y^{\rho}-\rho X)X|Y^{\rho}]\right)$$

$$= f_{Y^{\rho}}(Y^{\rho}) \left(X \mathbb{E}[T'(Z)|Y^{\rho}] - \mathbb{E}[T'(Z)X|Y^{\rho}]\right).$$

I-MMSE De Bruijn's identity

$$= \int_{\mathbb{R}} e^{T(\rho(X-x)+Z)} T'(\rho(X-x)+Z)(X-x)\mu_{X}(dx)$$

$$= \int_{\mathbb{R}} f_{Y^{\rho}|X=x}(Y^{\rho}; x) T'(\rho(X-x)+Z)(X-x)\mu_{X}(dx)$$

$$= f_{Y^{\rho}}(Y^{\rho}) \int_{\mathbb{R}} T'(Y^{\rho}-\rho x)(X-x))\mu_{X|Y^{\rho}}(dx; Y^{\rho}) \text{ (explain later)}$$

$$= f_{Y^{\rho}}(Y^{\rho}) \left(X\mathbb{E}[T'(Y^{\rho}-\rho X)|Y^{\rho}] - \mathbb{E}[T'(Y^{\rho}-\rho X)X|Y^{\rho}]\right)$$

$$= f_{Y^{\rho}}(Y^{\rho}) \left(X\mathbb{E}[T'(Z)|Y^{\rho}] - \mathbb{E}[T'(Z)X|Y^{\rho}]\right).$$

$$f_{Y^{\rho}|X=x}(y; x)\mu_{X}(dx) = f_{Y^{\rho}|X=x}(y; x)f_{X}(x)dx = f_{XY^{\rho}}(x, y)dx$$

$$= f_{Y^{\rho}}(y)f_{X|Y^{\rho}=y}(x; y)dx = f_{Y^{\rho}}(y)\mu_{X|Y^{\rho}=y}(dx; y).$$

I-MMSE De Bruijn's identity

Hence we have

$$\begin{aligned} & \frac{d}{d\rho} I(X; Y^{\rho}) \\ &= -\frac{d}{d\rho} \mathbb{E}[\log f_{Y^{\rho}}(Y^{\rho})] = -\mathbb{E}\left[\frac{1}{f_{Y^{\rho}}(Y^{\rho})}\frac{d}{d\rho}f_{Y_{\rho}}(Y^{\rho})\right] \\ &= -\mathbb{E}\left(X\mathbb{E}[T'(Z)|Y^{\rho}]\right) + \mathbb{E}\left(\mathbb{E}[T'(Z)X|Y^{\rho}]\right) \\ &= -\mathbb{E}\left(X\mathbb{E}[T'(Z)|Y^{\rho}]\right) + \mathbb{E}[T'(Z)X] \\ &= -\mathbb{E}\left(X\mathbb{E}[T'(Z)|Y^{\rho}]\right). \end{aligned}$$

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I-MMSE De Bruijn's identity

• When $T(x) = -x^2/2$,

 $-\mathbb{E}(X\mathbb{E}[T'(Z)|Y^{\rho}]) = \mathbb{E}(X\mathbb{E}[Z|Y^{\rho}]) = \mathbb{E}(X\mathbb{E}[Y^{\rho} - \rho X|Y^{\rho}])$

I-MMSE De Bruijn's identity

• When $T(x) = -x^2/2$,

 $-\mathbb{E}(X\mathbb{E}[T'(Z)|Y^{\rho}]) = \mathbb{E}(X\mathbb{E}[Z|Y^{\rho}]) = \mathbb{E}(X\mathbb{E}[Y^{\rho} - \rho X|Y^{\rho}])$ $= \mathbb{E}(XY^{\rho}) - \rho\mathbb{E}(X\mathbb{E}[X|Y^{\rho}]) = \rho\mathbb{E}(X^{2}) + \mathbb{E}(XZ) - \rho\mathbb{E}(X\mathbb{E}[X|Y^{\rho}])$

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• When $T(x) = -x^2/2$,

$$\begin{split} &-\mathbb{E}(X\mathbb{E}[T'(Z)|Y^{\rho}]) = \mathbb{E}(X\mathbb{E}[Z|Y^{\rho}]) = \mathbb{E}(X\mathbb{E}[Y^{\rho} - \rho X|Y^{\rho}]) \\ &= \mathbb{E}(XY^{\rho}) - \rho\mathbb{E}(X\mathbb{E}[X|Y^{\rho}]) = \rho\mathbb{E}(X^{2}) + \mathbb{E}(XZ) - \rho\mathbb{E}(X\mathbb{E}[X|Y^{\rho}]) \\ &= \rho\mathbb{E}(X^{2}) - \rho\mathbb{E}(X\mathbb{E}[X|Y^{\rho}]) = \rho\mathbb{E}(X - \mathbb{E}[X|Y^{\rho}])^{2}. \end{split}$$

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I-MMSE De Bruijn's identity

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I-MMSE De Bruijn's identity

Theorem

Let X be any random variable with a finite variance with density $f_X(x)$. Let Z be an independent standard normally distributed random variable. Then

$$\frac{d}{d\rho}H(X+\sqrt{\rho}Z)=\frac{1}{2}J(X+\sqrt{\rho}Z),$$

where $J(\cdot)$ is the Fisher information.

- Let $Y^{\rho} = X + \sqrt{\rho}Z$.
- Density function of Y^{ρ} :

$$f_{Y^{\rho}}(y) = \int_{\mathbb{R}} \frac{f_X(x)}{\sqrt{2\pi\rho}} \exp\left(-\frac{(y-x)^2}{2\rho}\right) dx$$

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I-MMSE De Bruijn's identity

$$f_{Y^{\rho}}(Y^{\rho}) = f_{Y^{\rho}}(X + \sqrt{\rho}Z) = \int_{\mathbb{R}} \frac{f_X(x)}{\sqrt{2\pi\rho}} \exp\left(-\frac{(X + \sqrt{\rho}Z - x)^2}{2\rho}\right) dx.$$

$$\begin{split} & \frac{d}{d\rho} f_{Y^{\rho}}(Y^{\rho}) \\ &= \int_{\mathbb{R}} \frac{f_X(x)}{\sqrt{2\pi\rho}} \exp\left(-\frac{(X+\sqrt{\rho}Z-x)^2}{2\rho}\right) \left(\frac{(X-x)(X+\sqrt{\rho}Z-x)}{2\rho^2} - \frac{1}{2\rho}\right) dx \\ &= f_{Y^{\rho}}(Y^{\rho}) \int_{\mathbb{R}} f_{X|Y^{\rho}}(x|Y^{\rho}) \left(\frac{(X-x)(Y^{\rho}-x)}{2\rho^2} - \frac{1}{2\rho}\right) dx. \end{split}$$

• The Blue terms = $f_{XY^{\rho}}(x, Y^{\rho})$.

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I-MMSE De Bruijn's identity

$$\begin{split} \frac{d}{d\rho} \mathcal{H}(Y^{\rho}) &= -\mathbb{E}\left[\frac{1}{f_{Y^{\rho}}(Y^{\rho})}\frac{d}{d\rho}f_{Y^{\rho}}(Y^{\rho})\right] \\ &= \mathbb{E}\left[\int_{\mathbb{R}} f_{X|Y^{\rho}}(x|Y^{\rho})\left(-\frac{(X-x)(Y^{\rho}-x)}{2\rho^{2}} + \frac{1}{2\rho}\right)dx\right] \\ &= \frac{1}{2\rho^{2}}\mathbb{E}\left(-XY^{\rho} + (X+Y^{\rho})\mathbb{E}[X|Y^{\rho}] - \mathbb{E}[X^{2}|Y^{\rho}]\right) + \frac{1}{2\rho} \\ &= \frac{1}{2\rho^{2}}\left(-\mathbb{E}[X^{2}] + \mathbb{E}[(\mathbb{E}[X|Y^{\rho}])^{2}]\right) + \frac{1}{2\rho} \end{split}$$

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I-MMSE De Bruijn's identity

• For the Fisher information
$$J(Y^{\rho}) = \mathbb{E}\left[\left(\frac{f'_{Y^{\rho}}(Y^{\rho})}{f_{Y^{\rho}}(Y^{\rho})}\right)^{2}\right]$$
.

$$f'_{Y\rho}(y) = \int_{\mathbb{R}} \frac{f_X(x)}{\sqrt{2\pi\rho}} \exp\left(\frac{-(y-x)^2}{2\rho}\right) \frac{x-y}{\rho} dx$$
$$= f_{Y\rho}(y) \int_{\mathbb{R}} f_{X|Y\rho}(x|y) \frac{x-y}{\rho} dx.$$

$$f_{Y^{\rho}}'(Y^{\rho}) = f_{Y^{\rho}}(Y^{\rho}) \int_{\mathbb{R}} f_{X|Y^{\rho}}(x|Y^{\rho}) \frac{x-Y^{\rho}}{\rho} dx = \frac{1}{\rho} f_{Y^{\rho}}(Y^{\rho}) \mathbb{E}\left[X-Y^{\rho}|Y^{\rho}\right].$$

I-MMSE De Bruijn's identity

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$$Y^{\rho} = X + \sqrt{\rho}Z$$

J

$$\begin{split} (Y^{\rho}) &= \mathbb{E}\left[\left(\frac{f_{Y^{\rho}}'(Y^{\rho})}{f_{Y^{\rho}}(Y^{\rho})}\right)^{2}\right] \\ &= \frac{1}{\rho^{2}}\mathbb{E}\left(\mathbb{E}\left[X - Y^{\rho}|Y^{\rho}\right]\right)^{2} = \frac{1}{\rho}\mathbb{E}\left(\mathbb{E}\left[Z|Y^{\rho}\right]\right)^{2} \\ &= \frac{1}{\rho^{2}}\left(\mathbb{E}\left[(\mathbb{E}[X|Y^{\rho}])^{2} + E\left[(Y^{\rho})^{2}\right] - 2\mathbb{E}[XY^{\rho}]\right), \end{split}$$

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$$\frac{d}{d\rho}H(Y^{\rho})=\frac{1}{2}J(Y^{\rho}),$$

since

$$\rho = \mathbb{E}[(X - Y^{\rho})^2].$$

Discrete-time Continuous-time

Theorem (Han-Song 2013)

Consider a discrete-time channel with feedback

$$Y_i = \rho g_i(X_i, Y_1^{i-1}) + Z_i, i = 1, \cdots, n,$$

where the density function of Z_i 's is $e^{T(x)}$. Then we have

$$\frac{dI(X_1^n; Y_1^n)}{d\rho} = \sum_{i=1}^n \left(-\mathbb{E}\left((g_i + \rho \frac{d}{d\rho} g_i) \mathbb{E}\left[\left[T'(Z_i) \right| Y_1^n \right] \right) + \mathbb{E}\left[(g_i + \rho \frac{d}{d\rho} g_i) \right] T'(Z_i) \right] \right)$$

where $g_i = g_i(X_i, Y_1^{i-1})$.

Discrete-time Continuous-time

• $f(y_1^n)$ means the density function of Y_1^n , $f(Y_1^n)$ means the density function of Y_1^n evaluated at Y_1^n , $f(y_1^n|x_1^n)$ means the conditional density function of Y_1^n given $X_1^n = x_1^n$, and so on.

$$I(X_1^n; Y_1^n) = H(Y_1^n) - H(Y_1^n | X_1^n) = -\mathbb{E} \log f(Y_1^n) - H(Y_1^n | X_1^n).$$

$$H(Y_1^n|X_1^n) = \sum_{i=1}^n H(Y_i|X_1^n, Y_1^{i-1}) = nH(Z_1).$$

Hence

$$\frac{dI(X;Y)}{d\rho} = -\mathbb{E}\left[\frac{d}{d\rho}\log f(Y_1^n)\right] = -\mathbb{E}\left[\frac{1}{f(Y_1^n)}\frac{d}{d\rho}f(Y_1^n)\right].$$

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Discrete-time Continuous-time

$$f(y_1^n|x_1^n) = f(y_1|x_1^n)f(y_2|y_1, x_1^n) \cdots f(y_n|y_1^{n-1}, x_1^n)$$

= $\prod_{i=1}^n \exp\left(T(y_i - \rho g_i(x_i, y_1^{i-1}))\right).$

$$\begin{aligned} &\frac{d}{d\rho}f(Y_1^n) \\ &= \frac{d}{d\rho}\int_{\mathbb{R}^n} f(Y_1^n|x_1^n)f(x_1^n)dx = \int_{\mathbb{R}^n} \frac{d}{d\rho}f(Y_1^n|x_1^n)f(x_1^n)dx \\ &= \int_{\mathbb{R}^n} \frac{d}{d\rho}\prod_{i=1}^n \exp\left(T\left(Y_i - \rho g_i(x_i, Y_1^{i-1})\right)\right)f(x_1^n)dx \\ &= \int_{\mathbb{R}^n} \frac{d}{d\rho}\prod_{i=1}^n \exp\left(T\left(\rho g_i(X_i, Y_1^{i-1}) - \rho g_i(x_i, Y_1^{i-1}) + Z_i\right)\right)f(x_1^n)dx \end{aligned}$$

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Discrete-time Continuous-time

$$= \int_{\mathbb{R}^n} f(Y_1^n | x_1^n) \sum_{i=1}^n T'(Y_i - \rho g_i(x_i, Y_1^{i-1}) \left[g_i(X_i, Y_1^{i-1}) - g_i(x_i, Y_1^{i-1}) \right. \\ \left. + \rho \frac{d}{d\rho} \left(g_i(X_i, Y_1^{i-1}) - g_i(x_i, Y_1^{i-1}) \right) \right] f(x_1^n) dx$$

$$= f(Y_1^n) \int_{\mathbb{R}^n} \sum_{i=1}^n T'(Y_i - \rho g(x_i, Y_1^{i-1}) \left[g_i(X_i, Y_1^{i-1}) - g_i(x_i, Y_1^{i-1}) \right. \\ \left. + \rho \frac{d}{d\rho} \left(g_i(X_i, Y_1^{i-1}) - g_i(x_i, Y_1^{i-1}) \right) \right] f(x_1^n | Y_1^n) dx.$$

• Denote $g_i := g_i(X_i, Y_1^{i-1})$ and $\tilde{g}_i := g_i(x_i, Y_1^{i-1})$.

Problem and background

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$$\begin{aligned} &\frac{d}{d\rho}f_X(Y_1^n) \\ =& f_X(Y_1^n) \sum_{i=1}^n \int_{\mathbb{R}^n} T'(Y_i - \rho \tilde{g}_i) \left[(g_i + \rho \frac{d}{d\rho}g_i) - (\tilde{g}_i + \rho \frac{d}{d\rho}\tilde{g}_i) \right] f_X(x_1^n | Y_1^n) dx \\ =& f_X(Y_1^n) \sum_{i=1}^n \left((g_i + \rho \frac{d}{d\rho}g_i) \mathbb{E} \left[T'(Y_i - \rho g_i) \right| Y_1^n \right] - \mathbb{E} \left[(g_i + \rho \frac{d}{d\rho}g_i) T'(Y_i - \rho g_i) \right| Y_1^n \right] \\ =& f_X(Y_1^n) \sum_{i=1}^n \left((g_i + \rho \frac{d}{d\rho}g_i) \mathbb{E} \left[T'(Z_i) \right| Y_1^n \right] - \mathbb{E} \left[(g_i + \rho \frac{d}{d\rho}g_i) T'(Z_i) \right| Y_1^n \right] \right). \end{aligned}$$

$$\frac{dI(X; Y^{\rho})}{d\rho} = -\mathbb{E}\left[\frac{d}{d\rho}\log f_X(Y_1^n)\right]$$
$$= \sum_{i=1}^n \left(-\mathbb{E}\left((g_i + \rho \frac{d}{d\rho}g_i)\mathbb{E}\left[T'(Z_i) \mid Y_1^n\right]\right) + \mathbb{E}\left[(g_i + \rho \frac{d}{d\rho}g_i)T'(Z_i)\right]\right).$$

Discrete-time Continuous-time

• For the Gaussian channel where $T(x) = -x^2/2$, we have

$$\begin{split} & \frac{dl(X; Y^{\rho})}{d\rho} \\ &= \sum_{i=1}^{n} \left(\mathbb{E} \left((g_{i} + \rho \frac{d}{d\rho} g_{i}) \mathbb{E} \left[Z_{i} | Y_{1}^{n} \right] \right) - \mathbb{E} \left[(g_{i} + \rho \frac{d}{d\rho} g_{i}) Z_{i} \right] \right) \\ &= \sum_{i=1}^{n} \left(\mathbb{E} \left((g_{i} + \rho \frac{d}{d\rho} g_{i}) (Y_{i} - \rho \mathbb{E} \left[g_{i} | Y_{1}^{n} \right] \right) \right) - \mathbb{E} \left[(g_{i} + \rho \frac{d}{d\rho} g_{i}) (Y_{i} - \rho g_{i}) \right] \right) \\ &= \sum_{i=1}^{n} \left(\rho \mathbb{E} \left[(g_{i} - \mathbb{E} (g_{i} | Y_{1}^{n}))^{2} \right] + \rho^{2} \mathbb{E} \left[(g_{i} - \mathbb{E} (g_{i} | Y_{1}^{n})) \frac{d}{d\rho} g_{i} \right] \right). \end{split}$$

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Discrete-time Continuous-time

Theorem (Han-Song 2013)

Consider the continuous-time channel with feedback

$$Y_t^{\rho} = \rho \int_0^t g(s, X_s, Y^{\rho}) ds + W_t, \ t \in [0, T],$$

$$(1)$$

where $g:[0,T] \times \mathbb{R} \times C([0,T];\mathbb{R}) \to \mathbb{R}$ is a bounded progressively measurable function. We have

$$\frac{dl(X; Y^{\rho})}{d\rho} = \rho \mathbb{E} \int_0^T (g_s - \mathbb{E}[g_s|Y^{\rho}])^2 ds + \rho^2 \int_0^T \mathbb{E} \left[(g_s - \mathbb{E}[g_s|Y^{\rho}]) \frac{\partial g_s}{\partial \rho} \right] ds,$$

where $g_s = g(s, X_s, Y^{\rho}).$

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Discrete-time Continuous-time

• For simplicity, consider the case $g(s, X_s, Y) = X_s$.

Lemma (Cameron-Martin)

 $Y_t = \rho \int_0^t X_s ds + W_t$, where W is a Wiener process, and X is independent of W. Then we have

$$\frac{d\mu_{Y|X}}{d\mu_W}(y;\omega) = \exp\left(\rho \int_0^T \omega_s dy_s - \frac{\rho^2}{2} \int_0^t \omega_s^2 ds\right), y \in C[0,T],$$

and

$$\frac{d\mu_Y}{d\mu_W}(y) = \int_{C[0,T]} \frac{d\mu_{Y|X}}{d\mu_W}(y;\omega)\mu_X(d\omega).$$

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Continuous-time

$$I(X; Y^{\rho}) = \mathbb{E}\left[\log \frac{d\mu_{XY^{\rho}}}{d(\mu_X \times \mu_{Y^{\rho}})}(X, Y^{\rho})\right] = \mathbb{E}\left[\log \frac{d\mu_{Y^{\rho}|X}}{d\mu_W}(Y^{\rho}; X)\right] - \mathbb{E}\left[\log \frac{d\mu_{Y^{\rho}}}{d\mu_W}(Y^{\rho})\right] = \frac{\rho^2}{2} \int_0^T \mathbb{E}[X_s^2] ds - \mathbb{E}\left[\log \frac{d\mu_{Y^{\rho}}}{d\mu_W}(Y^{\rho})\right]$$

$$\frac{d}{d\rho}I(X; Y^{\rho})$$
$$=\rho \int_{0}^{T} \mathbb{E}[X_{s}^{2}]ds - \frac{d}{d\rho}\mathbb{E}\left[\log\frac{d\mu_{Y^{\rho}}}{d\mu_{W}}(Y^{\rho})\right].$$

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Discrete-time Continuous-time

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$$\begin{split} & \frac{d}{d\rho} \left(\frac{d\mu_{Y\rho}}{d\mu_{W}} (Y^{\rho}) \right) \\ = & \frac{d}{d\rho} \int_{C[0,T]} \frac{d\mu_{Y\rho|X=\omega}}{d\mu_{W}} (Y^{\rho};\omega) \mu_{X}(d\omega) \\ = & \frac{d}{d\rho} \int_{C[0,T]} \exp\left\{ \rho \int_{0}^{T} \omega_{s} dY_{s}^{\rho} - \frac{\rho^{2}}{2} \int_{0}^{T} \omega_{s}^{2} ds \right\} \mu_{X}(d\omega) \\ = & \frac{d}{d\rho} \int_{C[0,T]} \exp\left\{ \rho^{2} \int_{0}^{T} \omega_{s} X_{s} ds + \rho \int_{0}^{T} \omega_{s} dW_{s} - \frac{\rho^{2}}{2} \int_{0}^{T} \omega_{s}^{2} ds \right\} \mu_{X}(d\omega) \\ = & \int_{C[0,T]} \left(\rho \int_{0}^{T} \omega_{s} (X_{s} - \omega_{s}) ds + \int_{0}^{T} \omega_{s} dY_{s}^{\rho} \right) \frac{d\mu_{Y^{\rho}|X=\omega}}{d\mu_{W}} (Y^{\rho};\omega) \mu_{X}(d\omega) \end{split}$$

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Discrete-time Continuous-time

$$= \int_{C[0,T]} \left(\rho \int_0^T \omega_s (X_s - \omega_s) ds + \int_0^T \omega_s dY_s^{\rho} \right) \frac{d\mu_{XY\rho}}{d\mu_W} (d\omega, Y^{\rho})$$

$$= \frac{d\mu_{Y\rho}}{d\mu_W} (Y^{\rho}) \int_{C[0,T]} \left(\rho \int_0^T \omega_s (X_s - \omega_s) ds + \int_0^T \omega_s dY_s^{\rho} \right) \mu_{X|Y\rho} (d\omega; Y^{\rho})$$

$$= \frac{d\mu_{Y\rho}}{d\mu_W} (Y^{\rho}) \left(\rho \int_0^T \mathbb{E} \left[X_s | Y^{\rho} \right] X_s ds - \rho \int_0^T \mathbb{E} \left[X_s^2 | Y^{\rho} \right] ds + \mathbb{E} \left[\int_0^T X_s dY_s^{\rho} | Y^{\rho} \right] \right)$$

Problem and background

Discrete-time Continuous-time

Comparison between old approach and our observation New proofs based on our observation Main results

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$$\begin{split} & \mathbb{E}\left[\frac{d}{d\rho}\left(\frac{d\mu_{Y^{\rho}}}{d\mu_{W}}(Y^{\rho})\right)\frac{d\mu_{W}}{d\mu_{Y^{\rho}}}(Y^{\rho})\right] \\ &=\rho\int_{0}^{T}\mathbb{E}\left[\mathbb{E}\left[X_{s}|Y^{\rho}\right]X_{s}\right]ds - \rho\int_{0}^{T}\mathbb{E}\left[X_{s}^{2}\right]ds + \mathbb{E}\left[\int_{0}^{T}X_{s}dY_{s}^{\rho}\right] \\ &=\rho\int_{0}^{T}\mathbb{E}\left[\left(\mathbb{E}\left[X_{s}|Y^{\rho}\right]\right)^{2}\right]ds. \end{split}$$

$$\frac{d}{d\rho}I(X; Y^{\rho})$$

$$=\rho \int_{0}^{T} \mathbb{E}[X_{s}^{2}]ds - \frac{d}{d\rho}\mathbb{E}\left[\log\frac{d\mu_{Y^{\rho}}}{d\mu_{W}}(Y^{\rho})\right]$$

$$=\rho \int_{0}^{T} \mathbb{E}\left[\left(X_{s} - \mathbb{E}\left[X_{s}|Y^{\rho}\right]\right)^{2}\right]ds.$$

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Further study

- I-MMSE when W is a general Gaussian process.
- Applications of I-MMSE.

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D. Guo, S. Shamai, and S. Verdu.

Mutual information and minimum mean-square error in Gaussian channels. IEEE Trans. Info. Theory, vol. 51, no. 4, pp. 1261-1282, 2005.

D. Guo, S. Shamai, and S. Verdu.

Additive non-Gaussian noise channels: mutual information and conditional mean estimation. IEEE ISIT, pp. 719-723, 2005.

S. Shamai.

From Constrained Signaling to Network Interference Via An Information-Estimation Perspective. Shannon Lecture, IEEE ISIT, 2011.

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THANK YOU!