Optimal Adaptive Strategies for Sequential Quantum Hypothesis Testing

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Outline

Setup

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- Quantum Hypothesis Testing: Literature Review
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• Binary hypothesis testing: $H_0: P = P_0$ and $H_1: P = P_1$, where P_0 and P_1 are probability distributions defined on the same alphabet \mathcal{X} .

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- T is the first time k that $d_k \neq *$.
- d_T is a {0,1}-valued function and d_T = i means H_i is the underlying hypothesis.
- $\alpha = P_0(d_T = 1)$ and $\beta = P_1(d_T = 0)$.

Error Exponents

Given a sequence of SHTs $\{\{(d_{n,k}, T_n)\}_{k=1}^{\infty}\}_{n=1}^{\infty}$ satisfying the constraint

 $\max_{i=0,1} \mathbf{E}_i[\mathcal{T}_n] \le n,$

the error exponents (E_0, E_1) defined as

$$E_0 = \liminf_{n \to \infty} \frac{1}{n} \log \frac{1}{\alpha_n}$$
 and $E_1 = \liminf_{n \to \infty} \frac{1}{n} \log \frac{1}{\beta_n}$.

In 1948, Wald and Wolfowitz showed that $E_0 \leq D(P_1 || P_0)$ and $E_1 \leq D(P_0 || P_1)$ and the error exponent can be achieved by a sequence of SPRTs.

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Setup

- \mathbb{C}^d : *d*-dimensional complex Euclidean space;
- Quantum state ρ over C^d: a d × d positive-semidefinite matrix with trace value 1.
- A positive operator-valued measure (POVM)
 m = {m(x) : x ∈ X} with outcomes in X: m(x) is a d × d
 positive-semidefinite matrix such that ∑_{x∈X} m(x) = I_d.
- A projector-valued measure m = {m(x) : x ∈ X} with outcomes in X: m is a POVM with the additional condition that m(x)² = m(x).
- The probability of obtaining outcome x when m is applied to the underlying state ρ:

$$P_{\rho,m}(X=x)=\mathrm{Tr}[\rho m(x)].$$

Quantum Divergence

For a $d \times d$ Hermitian matrix A, let $\sum_{i=1}^{d} \lambda_i P_i$ be the spectral decomposition of A. We define $\log A \triangleq \sum_{i=1}^{d} \log \lambda_i P_i$. For a $d \times d$ Hermitian matrix A, the *support* of A is defined as the subspace generated by the eigenvectors corresponding to non-0 eigenvalues.

Definition 1 (Quantum Relative Entropy)

Let ρ_0 and ρ_1 be two quantum states with the same support. The quantum relative entropy between two quantum states ρ_0 and ρ_1 is defined as

 $D(\rho_0 \| \rho_1) = \text{Tr}[\rho_0(\log \rho_0 - \log \rho_1)].$

The measured relative entropy is defined as

$$D_{\mathcal{M}}(\rho_0 \| \rho_1) := \sup_m D(P_{\rho_0, m} \| P_{\rho_1, m}), \tag{1}$$

where the supremum runs over all rank-1 PVMs comprised of d projectors.

Theorem 2 (Berta, Fawzi, Tomamichel (2017))

For two states ρ_0 and ρ_1 with full support, we have

$$D_{\mathcal{M}}(\rho_0 \| \rho_1) = \sup_{\mathcal{X}} \sup_{m \in \mathcal{M}_{\mathcal{X}}} D(P_{\rho_0, m} \| P_{\rho_1, m})$$
(2)

and the supremum is achieved at some PVM m^* with $|\mathcal{X}| = d$.

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- If the outcome is 0, guess the state $\rho_0^{\otimes n}$; otherwise guess the state $\rho_1^{\otimes n}$.
- Probability of type I-error: $\alpha_n := \text{Tr}[\rho_0^{\otimes n}(I_n \Lambda_n)].$

- The quantum system is prepared in one of the two states ρ₀^{⊗n} and ρ₁^{⊗n}.
- Make a quantum measurement $\{\Lambda_n, I_n \Lambda_n\}$ with outcome $\{0, 1\}$ to figure out the underlying states.
- If the outcome is 0, guess the state ρ₀^{⊗n}; otherwise guess the state ρ₁^{⊗n}.
- Probability of type I-error: $\alpha_n := \text{Tr}[\rho_0^{\otimes n}(I_n \Lambda_n)].$
- Probability of type II-error: $\beta_n := \text{Tr}[\rho_1^{\otimes n} \Lambda_n].$

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Quantum Stein's Lemma [Hiai and Petz (1991), Ogawa and Hayashi (2004)]

For any $\varepsilon > 0$, let $\beta_n(\varepsilon) \triangleq \inf \{\beta_n : (\Lambda_n, I_n - \Lambda_n) \text{ such that } \alpha_n \leq \varepsilon \}$. Then

$$\lim_{n\to\infty}\frac{1}{n}\log\frac{1}{\beta_n(\varepsilon)}=D(\rho_0\|\rho_1).$$

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Quantum Chernoeff exponent [Audenaert et al. (2007), Nussbaum and Azkoła (2009)]

Let
$$C_n = \inf \{ \alpha_n + \beta_n : (\Lambda_n, I_n - \Lambda_n) \}$$
. Then

$$\lim_{n\to\infty}\frac{1}{n}\log\frac{1}{C_n}=C(\rho_0,\rho_1)=-\log\sup_{0\leq s\leq 1}\mathrm{Tr}[\rho_0^s\rho_1^{1-s}].$$

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Quantum Hoeffding's Bound [Hayashi and Nagaoka (2003), Ogawa and Hayashi (2004), Nagaoka (2006)]

Let

$$\mathcal{R} := \left\{ (R_0, R_1) : \exists (\Lambda_n, I_n - \Lambda_n) \text{ such that } \begin{array}{l} R_0 \leq \liminf_{n \to \infty} \frac{1}{n} \log \frac{1}{\alpha_n}, \\ R_1 \leq \liminf_{n \to \infty} \frac{1}{n} \log \frac{1}{\beta_n}. \end{array} \right\}$$

Then

$$\mathcal{R} = \bigcup_{0 \le r \le D(\rho_1 \| \rho_0)} \left\{ (R_0, R_1) : \begin{array}{c} R_0 \le r, \\ R_1 \le \sup_{0 \le s \le 1} \frac{-sr - \log \operatorname{Tr}[\rho_0^{1-s} \rho_1^s]}{1-s} \end{array} \right\}.$$

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Figure: The structure of a general adaptive sequential hypothesis testing protocol.

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Sequential Quantum Hypothesis Test

A sequential quantum hypothesis test (SQHT) $S = (\mathcal{X}, \{p_k, d_k\}_{k=1}^{\infty})$, in its most general form, is given by (see also Figure 1):

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- a sequence of (conditional) probability distributions to determine the next measurement, p_k(m_k|x₁^{k-1}, m₁^{k-1}) for every k ∈ N;
- a sequence of $\{0, 1, *\}$ -valued decision functions $d_k(x_1^k, m_1^k)$, for every $k \in \mathbb{N}$.

Let T be the first time that $d_k \neq *$. Thus the number of samples of the underlying state ρ used during the test is T.

Sequential Quantum Probability Ratio Tests (SQPRTs)

For $k \ge 1$, let p_k be the probability density function according to which the experimenter chooses POVM M_k at time k. Let

$$S_k := \log \frac{\mathbb{P}_0(X_1^k, M_1^k)}{\mathbb{P}_1(X_1^k, M_1^k)}.$$
(3)

Additionally, let A and B be two fixed positive real numbers. The decision function d_k at time k is defined as follows

$$d_k(X_1^k, M_1^k) = egin{cases} 0 & S_k \geq B \ 1 & S_k \leq -A \ st & ext{otherwise.} \end{cases}$$

Let $T = \inf\{k \ge 1 : S_k \notin (-A, B)\}$. We say $(\mathcal{X}, \{p_k, d_k\}_{k=1}^{\infty}, T)$ is an *SQPRT* with parameters A and B.

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Adaptive Versus Non-adaptive Strategies

- In an adaptive QSHT S, the measurement m_{k+1} depends on m_1^k and x_1^k .
- In an non-adaptive QSHT S, the measurement m_{k+1} does not depends on m_1^k and x_1^k .

Consider sequences of QSHT S_n , indexed by $n \in \mathbb{N}$. We use $\mathbb{P}_{n,i}$ and $\mathbf{E}_{n,i}[\cdot]$ to denote the probability measure and the expectation induced by the QSHT S_n and the underlying state ρ_i .

Two Types of Errors

• Type I-error: $\alpha_n := \mathbb{P}_{n,0}(d_{\mathcal{T}_n} = 1).$

• Type II-error:
$$\beta_n := \mathbb{P}_{n,1}(d_{\mathcal{T}_n} = 0).$$

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• Type I-error:
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• Type II-error:
$$\beta_n := \mathbb{P}_{n,1}(d_{T_n} = 0).$$

Constraints on the Sample Size

• The first type of constraint is the expectation constraint:

$$\max_{i\in\{0,1\}}\mathbf{E}_{n,i}[T_n] \le n.$$
(5)

• The second type of constraint is the probabilistic constraint

$$\max_{i \in \{0,1\}} \mathbb{P}_{n,i}(T_n > n) < \varepsilon. \tag{6}$$

Achievable Error Exponent Pairs

A pair $(R_0, R_1) \in \mathbb{R}^2_+$ is said to be an achievable error exponent pair under the expectation constraint if there exists a sequence of QSHTs $\{S_n\}_{n\in\mathbb{N}}$ such that

$$\liminf_{n \to \infty} \frac{1}{n} \log \frac{1}{\alpha_n} \ge R_0, \text{ and } \liminf_{n \to \infty} \frac{1}{n} \log \frac{1}{\beta_n} \ge R_1, \text{ and } (7)$$
$$\limsup_{n \to \infty} \left(\max_{n \to \infty} \mathbf{E}_n \mid T \mid n \right) \le 0$$
(8)

$$\limsup_{n \to \infty} \left(\max_{i \in \{0,1\}} \mathbf{E}_{n,i}[\mathcal{T}_n] - n \right) \le 0.$$
(8)

Similarly, for $0 < \varepsilon < 1$, a pair $(R_0, R_1) \in \mathbb{R}^2_+$ is said to be an ε -achievable error exponent pair under the probabilistic constraint if there exists a sequence of QSHTs $\{S_n\}_{n\in\mathbb{N}}$ such that (7) hold and (instead of(8)),

$$\limsup_{n \to \infty} \max_{i \in \{0,1\}} \mathbb{P}_{n,i}(T_n > n) < \varepsilon.$$
(9)

Error Exponent Regions

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- $\mathcal{A}_{\rm E}$ is the closure of the set of all achievable error exponent pairs under the expectation constraint using adaptive strategies.
- A_P(ε) is the closure of the set of all ε-achievable error exponent pairs under the probabilistic constraint using adaptive strategies.
- $\mathcal{R}_{\rm E}$ is the closure of the set of achievable error exponent pairs under the expectation constraint using non-adaptive strategies.

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- $\mathcal{R}_{P}(\varepsilon)$ is the closure of the set of achievable error exponent pairs under the probabilistic constraint using non-adaptive strategies.

Sequential Quantum Hypothesis Testing: Literature Review

- Slussarenko et al. (2017, PRL) first applied sequential strategy to discriminate two quantum states.
- Martínez-Vargas et al. (2021, PRL) proposed the problem of sequential quantum hypothesis testing and showed the following.
 - For vanishing error probabilities α and $\beta\text{,}$

$$\begin{split} \mathbf{E}_0[T] &= \frac{1 + o(1)}{D(P_{\rho_0,m} \| P_{\rho_1,m})} \log \frac{1}{\beta}, \\ \mathbf{E}_1[T] &= \frac{1 + o(1)}{D(P_{\rho_1,m} \| P_{\rho_0,m})} \log \frac{1}{\alpha}. \end{split}$$

• For vanishing error probabilities α and $\beta,$

$$egin{aligned} \mathbf{E}_0[\mathcal{T}] &\geq rac{1+o(1)}{D(
ho_0\|
ho_1)}\lograc{1}{eta}, \ \mathbf{E}_1[\mathcal{T}] &\geq rac{1+o(1)}{D(
ho_1\|
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Main Results for Adaptive Strategies

Theorem 3

Let ρ_0 and ρ_1 be two quantum states with full support. Then for any $0 < \varepsilon < 1$,

$$\mathcal{A}_{\mathrm{P}}(\varepsilon) = \mathcal{A}_{\mathrm{E}} = \left\{ (R_0, R_1) : \begin{array}{c} R_0 \leq D_{\mathcal{M}}(\rho_1 \| \rho_0) \\ R_1 \leq D_{\mathcal{M}}(\rho_0 \| \rho_1) \end{array} \right\}.$$
(10)

Comparison with the result in Martínez-Vargas et al. (2021, PRL)

For vanishing error probabilities α and β ,

$$\begin{split} \mathbf{E}_0[T] &= \frac{1 + o(1)}{D_{\mathcal{M}}(\rho_1 \| \rho_0)} \log \frac{1}{\beta}, \\ \mathbf{E}_1[T] &= \frac{1 + o(1)}{D_{\mathcal{M}}(\rho_0 \| \rho_1)} \log \frac{1}{\alpha}. \end{split}$$

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Main Results for Adaptive Strategies



Figure: Multiple copies

Main Results for Adaptive Strategies



Figure: Multiple copies

Consider the binary quantum hypothesis test,

$$H_0^{(I)}: \rho^{\otimes I} = \rho_0^{\otimes I} \qquad H_1^{(I)}: \rho^{\otimes I} = \rho_1^{\otimes I}.$$
(11)

Define the achievable regions of the error exponent pairs $\mathcal{A}_{\rm E}^{(l)}$ and $\mathcal{A}_{\rm P}^{(l)}(\varepsilon)$. Similar to Theorems 3, we have

$$\mathcal{A}_{\rm E}^{(l)} = \mathcal{A}_{\rm P}^{(l)}(\varepsilon) = \left\{ (R_0, R_1) : \begin{array}{c} R_0 \leq \frac{1}{l} D_{\mathcal{M}}(\rho_1^{\otimes l} \| \rho_0^{\otimes l}) \\ R_1 \leq \frac{1}{l} D_{\mathcal{M}}(\rho_0^{\otimes l} \| \rho_1^{\otimes l}) \end{array} \right\}.$$
(12)

Main Results for Non-Adaptive Strategies: Ultimate Quantum Limit

Hiai and Petz, 1991

Let ρ_0 and ρ_1 be two quantum states with full support. Then

$$\lim_{l \to \infty} \frac{D_{\mathcal{M}}(\rho_1^{\otimes l} \| \rho_0^{\otimes l})}{l} = D(\rho_1 \| \rho_0).$$
(13)

We now characterize the ultimate quantum limit of achievable error exponent pairs using sequential adaptive testing strategies.

Theorem 4

Let ρ_0 and ρ_1 be two quantum states with full support. Then for any $0 < \varepsilon < 1$,

$$\bigcup_{l=1}^{\infty} \mathcal{A}_{\mathrm{E}}^{(l)} = \bigcup_{l=1}^{\infty} \mathcal{A}_{\mathrm{P}}^{(l)}(\varepsilon) = \left\{ (R_0, R_1) : \begin{array}{c} R_0 \leq D(\rho_1 \| \rho_0) \\ R_1 \leq D(\rho_0 \| \rho_1) \end{array} \right\}. \quad (14)$$



Figure: Schematic of the optimal trade-off between the exponential decay rates for the error of the first and second kind

Main Results for Non-Adaptive Strategies

Theorem 5

Let ρ_0 and ρ_1 be two quantum states with full support. Then for any $0 < \varepsilon < 1$,

$$\mathcal{R}_{\mathrm{E}} = \mathcal{R}_{\mathrm{P}}(\varepsilon) = \overline{\mathrm{Conv}(\mathcal{C})},$$
 (15)

where

$$\mathcal{C} = \bigcup_{\mathcal{X}} \bigcup_{m \in \mathcal{M}_{\mathcal{X}}} \left\{ (R_0, R_1) : \begin{array}{c} R_0 \leq D(P_{\rho_1, m} \| P_{\rho_0, m}) \\ R_1 \leq D(P_{\rho_0, m} \| P_{\rho_1, m}) \end{array} \right\}, \quad (16)$$

and \mathcal{X} runs over all finite sets and $\mathcal{M}_{\mathcal{X}}$ is the set of POVMs with support \mathcal{X} .

Main Results for Non-Adaptive Strategies

Corollary 6

Let ρ_0 and ρ_1 be two quantum states with full support and let

$$\mathcal{C}^{(I)} = \bigcup_{\mathcal{X}} \bigcup_{m \in \mathcal{M}_{\mathcal{X}}^{(I)}} \left\{ (R_0, R_1) : \begin{array}{c} R_0 \leq \frac{1}{I} D(P_{\rho_1, m} \| P_{0, m}) \\ R_1 \leq \frac{1}{I} D(P_{\rho_0, m} \| P_{1, m}) \end{array} \right\}.$$
(17)

Then for any $0 < \varepsilon < 1$, we have

$$\mathcal{R}_{\rm E}^{(l)} = \mathcal{R}_{\rm P}^{(l)}(\varepsilon) = \overline{\operatorname{Conv}(\mathcal{C}^{(l)})},$$
 (18)

and

$$\bigcup_{l=1}^{\infty} \mathcal{R}_{\rm E}^{(l)} = \bigcup_{l=1}^{\infty} \mathcal{R}_{\rm P}^{(l)}(\varepsilon).$$
(19)

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Let $\rho_0 = r_0 |\psi_0\rangle \langle \psi_0| + (1 - r_0)_2^I$ and $\rho_1 = r_1 |\psi_1\rangle \langle \psi_1| + (1 - r_1)_2^I$, where $|\psi_i\rangle = \cos \frac{\theta}{4} |0\rangle + (-1)^i \sin \frac{\theta}{4} |1\rangle$, $0 \le \theta \le \pi$, and $0 \le r_i \le 1$, $|0\rangle = (1,0)^{\top}$, $|1\rangle = (0,1)^{\top}$, I is the 2 × 2 identity matrix. For ρ_0 and ρ_1 with parameters (r_0, r_1, θ) , we define the sum rate of error exponent pairs as follows:

$$f(r_0, r_1, \theta) := D_{\mathcal{M}}(\rho_1 \| \rho_0) + D_{\mathcal{M}}(\rho_0 \| \rho_1)$$
(20)

and

$$g(r_0, r_1, \theta) := \sup_{\mathcal{X}} \sup_{m \in \mathcal{M}_{\mathcal{X}}} D(P_{\rho_0, m} \| P_{\rho_1, m}) + D(P_{\rho_1, m} \| P_{\rho_0, m}).$$
(21)

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Let
$$[d^2] = \{1, 2, \dots, d^2\}$$
 and let
 $\mathcal{M}^{(1)}_{[d^2]} = \{m \in \mathcal{M}_{[d^2]} : m(x) \text{ is of rank one for all } x \in [d^2]\}$

Theorem 7

Let ρ_0 and ρ_1 be two quantum states with full support. Then

$$\overline{\operatorname{Conv}(\mathcal{C})} = \overline{\operatorname{Conv}(\mathcal{C}^{(1)})}$$
(22)

where

$$\mathcal{C}^{(1)} = \bigcup_{m \in \mathcal{M}^{(1)}_{[d^2]}} \left\{ (R_0, R_1) : \begin{array}{l} R_0 \leq D(P_{\rho_1, m} \| P_{\rho_0, m}) \\ R_1 \leq D(P_{\rho_0, m} \| P_{\rho_1, m}) \end{array} \right\}.$$
(23)

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Figure: Maxima of the sum rate of error exponent pairs with adaptive or non-adaptive measurement strategies for $r_1 = r_2 = 0.98$ and $\theta \in (0, \frac{\pi}{2})$. The gap is most pronounced for large values of θ .



Figure: Achievable regions of error exponent pairs with adaptive or non-adaptive measurement strategies for when $(r_1, r_2, \theta) = (0.98, 0.98, 1.57)$. Note that the region for adaptive strategies is the entire rectangle including the region for non-adaptive strategies.

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Figure: Error probabilities for different tests when $(r_1, r_2, \theta) = (0.9, 0.9, \frac{\pi}{2}).$

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Main Idea: Construct an appropriate sequence of SQPRTs.

SQPRTs

Let $\mathcal{X} = \{1, 2, ..., d\}$. Let $m_0^* = \{m_0^*(x)\}_{x \in \mathcal{X}}$ and $m_1^* = \{m_1^*(x)\}_{x \in \mathcal{X}}$ be the projective-valued measurements that achieve the suprema in the definitions of $D_{\mathcal{M}}(\rho_0 \| \rho_1)$ and $D_{\mathcal{M}}(\rho_1 \| \rho_0)$, respectively. Recall that $Z_j = \log \operatorname{Tr} \left[\rho_0 M_j(X_j) \right] - \log \operatorname{Tr} \left[\rho_1 M_j(X_j) \right]$ and $S_k = \sum_{j=1}^k Z_j$.

Adaptive Strategies

For $k \ge 1$, the adaptive strategies are defined as follows

$$p_{k}(m_{0}^{*}|x_{1}^{k-1},m_{1}^{k-1}) = \begin{cases} \frac{1}{2} & \text{if } k = 1\\ 1 & \text{if } k \ge 2 \text{ and } S_{k-1} \ge 0, \end{cases} \quad \text{and} \quad (24)$$

$$p_{k}(m_{1}^{*}|x_{1}^{k-1},m_{1}^{k-1}) = \begin{cases} \frac{1}{2} & \text{if } k = 1\\ 1 & \text{if } k \ge 2 \text{ and } S_{k-1} < 0. \end{cases} \quad (25)$$

SQPRTs

For any fixed $0 < \tau < \min\{D_{\mathcal{M}}(\rho_{1} \| \rho_{0}), D_{\mathcal{M}}(\rho_{0} \| \rho_{1})\}$, let $A_{n} := n(D_{\mathcal{M}}(\rho_{1} \| \rho_{0}) - \tau)$ and $B_{n} := n(D_{\mathcal{M}}(\rho_{0} \| \rho_{1}) - \tau)$. (26) Let $T_{n} = \inf\{k \ge 1 : S_{k} \notin (-A_{n}, B_{n})\}$ and that $d_{n,k}(X_{1}^{k}, M_{1}^{k}) = \begin{cases} 0 & S_{k} \ge B_{n} \\ 1 & S_{k} \le -A_{n} \\ * & \text{otherwise.} \end{cases}$ (27)

We will show that this sequence of SQPRTs $S_n = (\mathcal{X}, \{p_k, d_{n,k}\}_{k=1}^{\infty}, T_n)$ with parameters A_n and B_n achieves $(D_{\mathcal{M}}(\rho_1 \| \rho_0), D_{\mathcal{M}}(\rho_0 \| \rho_1)).$

Bounds on the Error Probabilities

$$egin{aligned} & x_n = \mathbb{P}_0(d_{\mathcal{T}_n} = 1) \ & = \mathbf{E}_0[\chi_{\{S_{\mathcal{T}_n} \leq -A_n\}}] \ & = \mathbf{E}_1[e^{S_{\mathcal{T}_n}}\chi_{\{S_{\mathcal{T}_n} \leq -A_n\}}] \ & \leq e^{-A_n}. \end{aligned}$$

Similarly, $\beta_n \leq e^{-B_n}$.

Verification of the Expectation Constraints for the Tests S_n

Let $\hat{T}_n = \inf\{k : S_k \ge B_n\}$. Then $T_n \le \hat{T}_n$ and

$$S_{\hat{T}_n} \leq S_{\hat{T}_n-1} + Z_n \leq B_n + C.$$

$$\begin{split} \mathbf{E}_{0}[\mathcal{T}_{n}] &\leq \mathbf{E}_{0}[\hat{\mathcal{T}}_{n}] \\ &= \frac{-\mathbf{E}_{0}[S_{\hat{\mathcal{T}}_{n}} - \hat{\mathcal{T}}_{n}D_{\mathcal{M}}(\rho_{0}\|\rho_{1})] + \mathbf{E}_{0}[S_{\hat{\mathcal{T}}_{n}}]}{D_{\mathcal{M}}(\rho_{0}\|\rho_{1})} \\ &\leq \frac{-\mathbf{E}_{0}[S_{\hat{\mathcal{T}}_{n}} - \hat{\mathcal{T}}_{n}D_{\mathcal{M}}(\rho_{0}\|\rho_{1})] + B_{n} + C}{D_{\mathcal{M}}(\rho_{0}\|\rho_{1})} \\ &\leq \frac{C_{1} + B_{n} + C}{D_{\mathcal{M}}(\rho_{0}\|\rho_{1})} = n - \frac{n\tau - C_{1} - C}{D_{\mathcal{M}}(\rho_{0}\|\rho_{1})}. \end{split}$$

The following lemma provides lower bounds on the error probabilities for a general SQHT $(\mathcal{X}, \{p_k, d_k\}_{k=1}^{\infty}, T)$.

Lemma 8

For any SQHT $(\mathcal{X}, \{p_k, d_k\}_{k=1}^{\infty}, T)$ with adaptive strategies such that

$$\max_{i=0,1} \mathbf{E}_i[\mathcal{T}] < \infty, \tag{28}$$

the following inequalities hold,

$$\log \frac{1}{\beta} \leq \frac{\mathbf{E}_{0}[T]D_{\mathcal{M}}(\rho_{0}\|\rho_{1}) + 1}{1 - \alpha} \quad and \qquad (29)$$
$$\log \frac{1}{\alpha} \leq \frac{\mathbf{E}_{0}[T]D_{\mathcal{M}}(\rho_{0}\|\rho_{1}) + 1}{1 - \beta}. \qquad (30)$$

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Let $\{S_n\}_{n=1}^{\infty}$ be a sequence of SQHTs with adaptive strategies such that $\alpha_n \to 0$ and $\beta_n \to 0$ and the sequence $\{T_n\}_{n=1}^{\infty}$ satisfies the expectation constraint (8). Then from (29) and (30) in Lemma 8, we have that

$$\limsup_{n \to \infty} \frac{1}{n} \log \frac{1}{\beta_n} \le \limsup_{n \to \infty} \frac{\mathsf{E}_{n,0}[\mathcal{T}_n] D_{\mathcal{M}}(\rho_0 \| \rho_1) + 1}{n(1 - \alpha_n)} \le D_{\mathcal{M}}(\rho_0 \| \rho_1).$$
(31)

and

$$\limsup_{n \to \infty} \frac{1}{n} \log \frac{1}{\alpha_n} \le \limsup_{n \to \infty} \frac{\mathsf{E}_{n,1}[T_n] D_{\mathcal{M}}(\rho_1 \| \rho_0) + 1}{n(1 - \beta_n)} \le D_{\mathcal{M}}(\rho_1 \| \rho_0).$$
(32)

Conclusion: any achievable error exponent pair (R_0, R_1) is such that $R_0 \leq D_{\mathcal{M}}(\rho_1 \| \rho_0)$ and $R_1 \leq D_{\mathcal{M}}(\rho_0 \| \rho_1)$.

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Proof of Lemma 8

- Let P_i be the probability measure on (Ω, F) when the underlying state is ρ_i.
- Let *F_T* be the sub-*σ*-algebra generated by *T* and let P_{i,T} be the restriction of P_i to the *σ*-algebra *F_T*.

Then

$$\exp(S_{\mathcal{T}}) = rac{\mathrm{d}\mathbb{P}_{0,\mathcal{T}}}{\mathrm{d}\mathbb{P}_{1,\mathcal{T}}} ext{ and } \mathbf{E}_0[S_{\mathcal{T}}] = D(\mathbb{P}_{0,\mathcal{T}} \| \mathbb{P}_{1,\mathcal{T}}).$$

We define a stochastic kernel V with input alphabet Ω (with elements ω) and output alphabet $\{0,1\}$ as follows:

$$V(0|\omega) := \begin{cases} 1 & \text{if } d_T(\omega) = 0\\ 0 & \text{if } d_T(\omega) = 1 \end{cases}$$
(33)

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Let
$$D(a||b) = a \log \frac{a}{b} + (1-a) \log \frac{1-a}{1-b}$$
 for any $0 \le a, b \le 1$.

Proof of Lemma 8

Using the data processing inequality to the classical relative entropy when $(\mathbb{P}_{0,T}, \mathbb{P}_{1,T})$ is processed via the stochastic kernel V, we obtain,

$$D(\alpha || 1 - \beta) = D(\mathbb{P}_0(d_T = 1) || \mathbb{P}_1(d_T = 1))$$

$$\leq D(\mathbb{P}_{0,T} || \mathbb{P}_{1,T}) = \mathbf{E}_0[S_T], \qquad (34)$$

which implies that

$$\log \frac{1}{\beta} \le \frac{D(\alpha \| 1 - \beta) + 1}{1 - \alpha} \le \frac{\mathsf{E}_0[S_T] + 1}{1 - \alpha}.$$
 (35)

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Proof of Lemma 8

Key Fact: $\{S_k - kD_{\mathcal{M}}(\rho_0 \| \rho_1)\}_{k=1}^{\infty}$ is a supermartingale. Applying Optional Stopping Theorem to the supermartingale $\{S_k - kD_{\mathcal{M}}(\rho_0 \| \rho_1)\}_{k=1}^{\infty}$ and the stopping time *T*, we obtain

$$\mathbf{E}_{0}[S_{T} - TD_{\mathcal{M}}(\rho_{0} \| \rho_{1})] \le \mathbf{E}_{0}[S_{1} - D_{\mathcal{M}}(\rho_{0} \| \rho_{1})] \le 0.$$
(36)

Combining (35) and (36), we obtain

$$\log \frac{1}{\beta} \leq \frac{\mathsf{E}_0[S_T] + 1}{1 - \alpha} \leq \frac{\mathsf{E}_0[T]D_{\mathcal{M}}(\rho_0 \| \rho_1) + 1}{1 - \alpha}.$$
 (37)

Similarly, we have that

$$\log \frac{1}{\alpha} \le \frac{\mathbf{E}_0[\mathcal{T}]D_{\mathcal{M}}(\rho_0 \| \rho_1) + 1}{1 - \beta}.$$
(38)

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Lemma 9

For any $0 \le \Lambda_n \le I_n$, we have that $\operatorname{Tr}[\rho_0^{\otimes n}\Lambda_n] - \gamma \operatorname{Tr}[\rho_1^{\otimes n}\Lambda_n] \le \operatorname{Tr}[\rho_0^{\otimes n}\{\rho_0^{\otimes n} \ge \gamma \rho_1^{\otimes n}\}].$

Note that

$$P_0(T_n > k) = 1 - P_0(T_n \le n)$$

= $1 - \sum_{j=1}^k (P_0(T_n = j, d_n = 0) + P_0(T_n = j, d_n = 1)).$

Let $\tilde{\alpha}_k$ and $\tilde{\beta}_k$ be the type-I and type-II error probabilities of the test: H_0 is true if and only if $\{T_n \leq k, \delta_n = 0\}$ holds.

Then we have that $\tilde{\beta}_k \leq \beta_n$ and

$$\tilde{\alpha}_{k} = P_{0}(T_{n} > k) + \sum_{j=1}^{k} P_{0}(T_{n} = j, d_{n} = 1)$$
$$\tilde{\beta}_{k} = \sum_{j=1}^{k} P_{0}(T_{n} = j, d_{n} = 0).$$

Hence

$$P_{0}(T_{n} > k) = \tilde{\alpha}_{k} - \sum_{j=1}^{k} P_{0}(T_{n} = j, d_{n} = 1) \geq \tilde{\alpha}_{k} - \alpha_{n}$$
$$\geq 1 - \gamma_{n} \tilde{\beta}_{k} - \operatorname{Tr}[\rho_{0}^{\otimes k} \{\rho_{0}^{\otimes k} \geq \gamma_{n} \rho_{1}^{\otimes k}\}] - \alpha_{n}$$
$$= \operatorname{Tr}[\rho_{0}^{\otimes k} \{\rho_{0}^{\otimes k} \leq \gamma_{n} \rho_{1}^{\otimes k}\}] - \gamma_{n} \beta_{n} - \alpha_{n}.$$
(39)

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Let

$$k_n = \frac{1}{D(\rho_0 \| \rho_1) + \tau} \log \frac{1}{\beta_n} \quad \text{and} \quad \gamma_n = e^{k_n (D(\rho_0 \| \rho_1) + \tau_1)}$$
(40)

with $\tau_1 > \tau > 0$. Then we have that

$$n \geq \mathbf{E}_{0}[T_{n}] \geq P_{0}(T_{n} > k_{n})k_{n}$$

$$\geq (\operatorname{Tr}[\rho_{0}^{\otimes k_{n}} \{\rho_{0}^{\otimes k_{n}} \leq \gamma_{n}\rho_{1}^{\otimes k_{n}}\}] - \gamma_{n}\beta_{n} - \alpha_{n})k_{n}$$

$$\sim (1 - \beta_{n}^{\frac{\tau_{1} - \tau}{D(\rho_{0}||\rho_{1}) + \tau}} - \alpha_{n})\frac{1}{D(\rho_{0}||\rho_{1}) + \tau}\log\frac{1}{\beta_{n}}.$$
 (41)

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Thanks for Your Attention!